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RESEARCH ARTICLE

DOES EARNINGS PER SHARES (EPS) FACTOR REALLY MATTER FOR PRICING STOCKS ON THE VIETNAMESE STOCK MARKET?

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ABSTRACT

This study examined the effect of income factor on stock returns on Ho Chi Minh stock exchange using data collected from 2006 to 2014. Results from OLS regression showed that, compair to the other commonly-documented factors (market, size, value and liquidity factors), the income factor had a less significant impact on the stock returns. This mean that investors on stock market did not use EPS as a important variable in their stock-pricing model. Grounded on the findings, we put forward some recommendations.

INTRODUCTION

Sharpe (1964) and Lintner (1965) are pioneers in calibrating the tradeoff between risks and returns of stocks with the CAPM (Capital Asset Pricing Model), in which the returns on a stock is dependent on its market risk (measured by the beta coefficient). Fama and French (1992) find that together with the market factor, firm size and the book-to-market value can help also to predict the stock's returns (called the three-factor model). Other studies have also found a number of different factors that also explain the stock returns, for examples, among others, momentum, earnings per share (EPS) and liquidity factor (Carhart, 1997; Amihud and Mendelson, 1986; Amihud, 2002; etc.). The list of empirical studies on determinants of stock returns is unexhaustedfor stock markets around the world (e.g., Acharya and Pedersen, 2005; Keene and Peterson, 2007; AmihudvàMendelson, 1986; EleswarapuvàReinganum, 1993; Petersen vàFialkowski, 1994; BrenmanvàSubrahmanyam, 1996; ...). Like other emerging countries, the question of what factors may determine stock returns on the Vietnamese stock market is quite interesting for both academics and practioners. Some studies have been done for the Fama-French three-factor model (E.g., Truong Dong Loc and Duong Thi Hoang Trang, 2014; Tran ThiHai Ly, 2010), while others examine also the impact of liquidity factor on stock returns (Nguyen AnhPhong, 2012; Vo and Batten, 2010; Tran et al., 2013). However, whether or not the income factor (EPS) – a common

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determinants of stock returns in empirical studies for stock markets of other countries – can help to predict stock returns in the stock market of Vietnam is still an open issue. Filling this gap of the literature, in this study we investigate the role of the income factor in explaining stock returns on the Ho chi minh stock exchange.

METHODOLOGY

Liquidity measure

Liquidity measure is defined following Amihud (2002). It reads

$$ILLIQ_{iy} = \frac{1}{D_{iy}} \sum_{t=1}^{D_{iy}} \frac{\mid R_{iyd} \mid}{VOLD_{ivyd}}$$

in which, D_{iy} is the number of trading days for stock i in the year; R_{iyd} represents returns of stock i in day d of year y and $VOLD_{ivyd}$ indicates the respective trading volumes of R_{iyd}

Portfolios and variables construction

Based on firm size, book-to-market-value ratio (BE/ME), liquidity ratio (LIQ) and earnings per share ratio (EP), we construct portfolios and variables following the approach by Fama and French (1993) and other previous studies (Datar et al. (1998), Batten and Vo (2010), Tran et al. (2014)). In this study, the time frame of portfolio reconstruction is 12 months. If a company's cap is less than or equal to 50% of the market cap for the entire market, then it is classified as a small group

(Portfolio S), and vice versa categorized into a big group (portfolio B). BE/ME ratio, liquidity (LIQ) and Earning Per Shares (EP) are sorted by the same manner: companies in the 30% lowest value group isclassified intoportfolio L (Low), 40% of stocks with subsequent values will be listed as portfolio M (Medium) and the remaining 30% of stocks with the highest values iscategorized asportfolio H (High).

Table 1.Portfolios constructions

| Size | BE/ME (Book-to-Market) | | | | | | |
|---------|----------------------------|-------------------|----------------------------|--|--|--|--|
| | L (30%) | M (40%) | H (30%) | | | | |
| S (50%) | SL | SM | SH | | | | |
| B (50%) | BL | BM | BH | | | | |
| | EP (Earnings Per Share) | | | | | | |
| | E _L (30%) | $E_{\rm M}(40\%)$ | $E_{\rm H}(30\%)$ | | | | |
| S (50%) | SE_L | SE_M | SE_H | | | | |
| B (50%) | BE_L | BE_M | BE_H | | | | |
| | LIQ (Liquidity) | | | | | | |
| | $L_L(30\%)$ | $L_{\rm M}(40\%)$ | $L_{\rm H}(30\%)$ | | | | |
| S (50%) | SL_{L} | SL_M | SL_H | | | | |
| B (50%) | BL_L | BL_M | BL_{H} | | | | |

Four variables comprising of size factor (SMB), value factor (HML), liquidity factor (LIQ) and income factor (EP) are calculated using the approach by Fama-French (1993): Size factor (SMB-small minus big): defined by average returns on small portfolios (S) minus average returns on big portfolios (B).

$$SMB = \frac{(SL - BL) + (SM - BM) + (SH - BH)}{3}$$

Value factor (HML-high minus low): defined by average returns on high-BE/ME portfolios (H)minus average returns on low-BE/ME portfolios (L).

$$HML = \frac{(SH - SL) + (BH - BL)}{2}$$

Income factor (EPS-earnings per share): defined by average returns on low-EPS portfolios (E_L)minus average returns on high-EPS portfolios (E_H)

$$EP = \frac{(SE_L - SE_H) + (BE_L - BE_H)}{2}$$

Liquidity factor (LIQ): defined by average returns on low-liquidity portfolios (L_L)minus average returns on high-liquidity portfolios (L_H)

$$LIQ = \frac{(SL_L - SL_H) + (BL_L - BL_H)}{2}$$

Empirical model

Whilea number of previous studies have shown the applicability of the Fama-French three-factor model (1952) for the Vietnamese stock market (E.g., Truong Dong Loc and

Table 2. Descriptive statistics for all variables

| Variables | Observations | Mean | Std. | Min | Max | P-Value ADF |
|-----------|--------------|---------|--------|----------|---------|-------------|
| MP | 409 | -0.0793 | 4.4105 | -16.3330 | 13.3249 | 0.000 |
| SMB | 409 | 0.0200 | 1.6558 | -6.8403 | 5.6220 | 0.000 |
| HTL | 409 | 0.4175 | 2.8216 | -12.9171 | 11.9081 | 0.000 |
| EP | 409 | -0.2769 | 2.3050 | -10.2362 | 17.3031 | 0.000 |
| LIQ | 409 | -0.0319 | 2.7211 | -10.9656 | 10.0019 | 0.000 |

Table 3. Correlation matrix among variables

| | MP | SMB | HML | EP | LIQ |
|-----|--------|--------|---------|---------|-------|
| MP | 1.000 | | | | |
| SMB | -0.321 | 1.000 | | | |
| HML | 0.154 | -0.167 | 1.000 | | |
| EP | 0.100 | 0.064 | -0.0309 | -0.0333 | 1.000 |
| LIQ | -0.297 | 0.260 | -0.206 | -0.2452 | 1.000 |

Table 4. Regression results from the three-factor model by Fama-French augmented with the liquidity and income factors for all portfolios

| | Firm siz | e | | | | BE/ | ME | | | |
|-----------|----------|------------|--------|--------------------|---------|-------------|----------|--------|-------------|--------|
| | | В | S | В | S | | BM | SM | BM | SM |
| Liquidity | | b | | t(b) | | | b | | t(b) | |
| | L_L | 0.947 | 0.886 | 31.950 | 41.670 | L_L | 0.899 | 0.881 | 39.600 | 32.820 |
| | L_{H} | 0.931 | 0.903 | 45.940 | 29.980 | L_{H} | 0.887 | 0.940 | 31.560 | 47.570 |
| | | S | | t(s) | | | S | | t(s) | |
| | L_{L} | -0.122 | 0.931 | -1.560 | 16.580 | L_{L} | 0.782 | 0.680 | 13.040 | 9.590 |
| | L_{H} | -0.203 | 1.012 | -3.800 | 12.740 | L_{H} | 0.533 | -0.154 | 7.190 | -2.950 |
| | | h | | t(h) | | | h | | t(h) | |
| | L_L | 0.355 | 0.489 | 8.050 | 15.470 | L_L | 0.606 | 0.290 | 17.950 | 7.280 |
| | L_{H} | 0.121 | 0.722 | 4.040 | 16.140 | L_{H} | 0.697 | 0.016 | 16.680 | 0.550 |
| | | ψ | | $t(\psi)$ | | | ψ | | $t(\psi)$ | |
| | L_L | 0.809 | 0.185 | 16.520 | 5.270 | $L_{\rm L}$ | 0.237 | 0.295 | 6.310 | 6.650 |
| | L_{H} | -0.327 | -0.679 | -9.760 | -13.660 | L_{H} | -0.534 | -0.303 | -11.510 | -9.300 |
| | | π | | $t(\mathcal{\pi})$ | | | π | | $t(\pi)$ | |
| | L_L | 0.236 | -0.009 | 4.370 | -0.230 | L_{L} | 0.032 | -0.021 | 0.770 | -0.440 |
| | L_{H} | 0.123 | 0.104 | 3.340 | 1.900 | L_{H} | 0.166 | 0.073 | 3.230 | 2.030 |
| | | α | | $t(\alpha)$ | | | α | | $t(\alpha)$ | |
| | L_L | -0.641 | -0.573 | -5.250 | -6.540 | L_L | -0.528 | -0.669 | -5.650 | -6.050 |
| | L_{H} | -0.563 | -0.651 | -6.750 | -5.250 | L_{H} | -0.522 | -0.642 | -4.510 | -7.890 |
| | | Adjusted-R | 2 | | | | | $-R^2$ | | |
| | L_L | 0.761 | 0.839 | | | L_L | 0.833 | 0.742 | | |
| | L_{H} | 0.889 | 0.822 | | | $L_{\rm H}$ | 0.832 | 0.890 | | |

^{*}Notes: L_L and L_Hare low and high liquidity, respectively. t(.) is t-value for regression coefficients.

Duong Thi Hoang Trang, 2014; Tran Thi Hai Ly, 2010), other research finds that liquidity factor is also an important factor for pricing stocks on the Vietnamese stock market (Nguyen AnhPhong, 2012; Vo and Batten, 2010; Tran *et al.*, 2013). In this study, we estimate a model with five factors, i.e. the Fama-French three-factor model (1952) augmented with theliquidity (LIQ)and Earning Per Shares (EP) factors, to test for the role of EP in pricing stocks. Specifically, the empirical model is shown as:

$$R_{pt} - R_{ft} = \alpha_p + b_p M P_t + s_p S M B_t + h_p H T M_t + \psi_p L I Q_1 + \pi_p E P_t + \varepsilon_{pt}$$

in which: R_{pt} is the average portfolio returns at time t; R_{ft} is risk-free rate at time t; MP represents the excess returns of the market (market returns minus risk-free rates at time t); SMB_t (small-minus-big) indicates size factor at time t; HML_t (high-minus-low) is value factor at time t; LIQ_t(low-minus-high liquidity) shows liquidity factor at time t; EP_t is income factor at time t and \mathcal{E}_{pt} is error terms. All portfolios are first sorted based on three criteria including firm size, BE/ME ratios and liquidity. Sorting by this manner can control for effects of size, value and liquidity factorson stock returns. Then regressions on stock portfolios are implemented using the OLS method. Variables are checked for stationary with ADF test. OLS

Data

All stock prices and market price index data are obtained from Ho Chi Minh Stock Exchange from 1/2006 to 12/2014.All data are in monthly frequency. Financial reports of listed firms are gathered from their websites.Government bond rates proxies for risk-free rates, collected from the stock market's website.

assumptions are also tested and corrected for violations using

Empirical results

Descriptive statistics

appropriate methods.

Table 2 shows the summary statistics for variables. In general, since no outliers in data can be observed, OLS estimations is appropriate. Besides, all variables are stationary. As can be seen from Table 3, the correlation between variables is fairly small (almost less than 0.3). Therefore, the possible effects of multi-collinearity in OLS regressions are negligible.

Regression results

Table 4 presents results of the OLS regression estimations for all portfolios sorted by firm size, BE/ME ratio and liquidity factor. As can be seen, all coefficients of the market factor (b) for all portfolios are statistically positive at 1% level and very close to 1.0. The coefficient of size factor (s) is statistically significant in 7 portfolios, only one coefficients are not statistically significant (i.e. big size portfolios (B) - liquidity is BL_L). For the value factor, all coefficients (h) for most portfolios are positive and statistical significance at the 1% level, except for SML_H. For the liquidity factor, all regression coefficients (ψ) are statistically significant at the 1% level. As expected, these results are similar to those in the previous studies (E.g., Lam and Tam (2011), Keene and Peterson (2007), Fama and French (1993)), confirming the applicability of our investigating approach. As for the income factor (EP), the

regression coefficient of only 4 portfolios is statistically positive at the 1% to 5% level, meanwhile that of the others are not (i.e. portfolios: SL_L, BML_L, SML_L). Noticeably, although significantly, all the coefficients are relatively smaller than those of other factors. Remarkably, all intercepts (α) for portfolios are significantly negative at about 1% level with the absolute value of below 0.6. To sum up, compare to the others comprising of market, size, value and liquidity factors, the income factor does not seem to have a significant impact on the stock-pricing model on the Vietnamese stock market. In other words, investors on the stock market do not take into account of the EPS ratio when valuing stocks. These results could be explained by the fact that stock market in developing countries such as Vietnam often suffer from poor accounting practices, price manipulation, and so forth. As a consequence, the market growth does not necessarily mean that they are sophisticated, and stock pricesis potentially driven by speculative motives (Shirai, 2004). In fact, it has been argued that emerging stock markets are highly volatile and may not reflect market fundamentals but instead be affected by "bubbles" or factors other than the present value of the expected future profits (Goergen and Renneboog, 2001, Bond et al., 2004). These reasons may result in an untrue value of EPS, this in turn deteriorates the market participants' beliefs in the information signaling by earnings-per-share ratio. Investors hence exclude the income factor from the stock-pricing model.

Conclusion and recommendation

This study examines the role of income factor (EPS) in pricing common stocks in the Ho Chi Minh Stock Exchange. Compare to the other factors comprising of market, size, value and liquidity, we find that the income factor seems to play a less important role in the stock-pricing model. In other words, investors may not consider the income factor when valuing shares in the Vietnam stock market. These results could be explained by the investors' beliefs in the accuracy of EPS ratio (and hence its contained information) due to poor accounting practices, price manipulation, high volatility and bubbles of stock prices. A number of recommendations for policy makers and stock market managers have been putforward. Firstly, stricter supervision rules should be imposed to avoid illegal stock prices manipulations leading to price bubbling. In addition, more transparency and timely information declaration to investors should also be taken care of to avoid the asymmetry, which can reduce speculative motives. Finally, higher standardized accounting standards should also be applied for listed firms.

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